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A banking explanation of the US velocity of money: 1919–2004

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ABSTRACT

The paper shows that US GDP velocity of M1 money has exhibited long cycles around a 1.25% per year upward trend, during the 1919–2004 period. It explains the velocity cycles through shocks constructed from a DSGE model and annual time series data (Ingram et al., 1994). Model velocity is stable along the balanced growth path, which features endogenous growth and decentralized banking that produces exchange credit. Positive shocks to credit productivity and money supply increase velocity, as money demand falls, while a positive goods productivity shock raises temporary output and velocity. The paper explains such velocity volatility at both business cycle and long run frequencies. With filtered velocity turning negative, starting during the 1930s and the 1987 crashes, and again around 2003, results suggest that the money and credit shocks appear to be more important for velocity during less stable times and the goods productivity shock more important during stable times.

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1. Introduction

Ireland (1994) modeled the interesting idea that the velocity of money could rise forever, until money was no longer being used, due to continuous technological progress that lowers the cost of credit. This is also the argument that Woodford (2003) makes for why it makes sense to consider monetary policy in an economy without money, the main paradigm of his highly influential work. And this is followed up by Gali (2008), in which money demand plays no role except to equilibrate the money supply as determined residually by the Taylor rule. Hromcova (2008) novelly reformulates the Ireland possibility of an ever rising velocity by using a human capital externality to lower the cost of credit.

Hromcova (2008) uses a Lucas (1988) type endogenous growth economy with human capital and an exchange technology that allows either cash or costly credit use, across a continuum of stores, similar to the cash-credit mix in Gillman (1993) and Ireland (1994). However, the cost of credit use is defined so that it falls continually as the human capital level in the economy increases. The consumer naturally chooses more credit use over time, and so velocity increases, and the balanced path equilibrium is defined as one in which velocity rises steadily, towards infinity.

We follow the same, well-respected, Lucas (1988) endogenous growth approach, which has support going back to Kocherlakota and Source (1996), however, in the form without any human capital externalities which are known to give rise to multiple or indeterminate equilibria. Our cost of credit is not postulated in a general transaction cost form, as in Hromcova (2008), Schmitt-Grohe and Uribe (2004), or as in Bansal and Coleman (1996). Rather it is derived from an equilibrium where credit is produced by an industry-based banking sector with constant returns to scale technology as in Clark (1984). With shocks to credit sector productivity, along with the money supply and goods sector productivity shocks as in the monetary real

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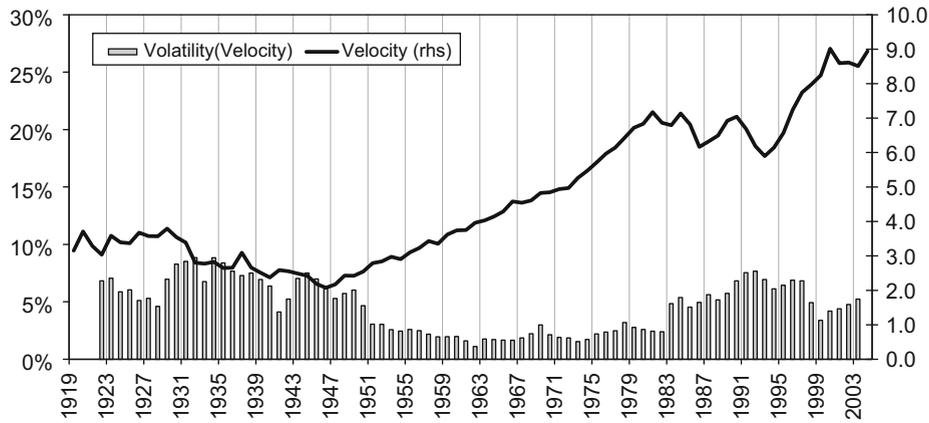


Fig. 1. Income velocity level and its volatility [Velocity = (nominal GDP/M1). Volatility is defined as the standard deviation over a 7-year moving window].

business cycle (RBC) approach, this economy has a well-defined balanced growth path (BGP) equilibrium with a stationary velocity, and not one rising forever (Section 2). Our contribution is then to use this baseline model to explain cycles of the US M1 velocity over a long historical period, with an extension to the UK money velocity for post-1978 data.

Simulations of the model show a good ability to replicate business cycle correlations, both real and “monetary”, while explaining 87% of the relative volatility of the M1 income velocity (Section 3). In Fig. 1 graph of US M1 income velocity and its volatility, the data show a 1.25% average annual upwards trend in the 1919–2004 velocity; in the analysis we use an 86 year Christiano and Fitzgerald (2003) filter to detrend the data minimally to achieve stationarity while leaving in all but the very long run trend component of the data. A decomposition of the velocity volatility provides estimates of the contribution to volatility by each of the three shocks, across various subperiods, and within both business cycle and long run windows. The shocks are constructed from actual annual time series data for the choice variables using the model’s equilibrium solution, following the methodology of Ingram et al. (1994). In this way the selected group of theoretically plausible RBC style shocks are “backed-out” of the model as in our previous work (Benk et al., 2005, 2008) and in Nolan and Thoenissen (2009), while constructed so as to have zero mean over the sample period. The paper also illustrates a graphical way to see how the shocks “add up” to explain the velocity cycles over the sample period, and discusses both volatility and level results for the US (Section 4), with extension to the UK (Section 5), and with literature comparison (Section 6).

2. The model

The representative agent economy is extended from the stochastic shock framework of Benk et al. (2008) by decentralizing the bank sector that produces credit; this decentralization follows that of Gillman and Kejak (2009), although there the economy is deterministic. By combining the business cycle with endogenous growth, stationary inflation lowers the output growth rate as supported empirically for example in Gillman et al. (2004) and Fountas et al. (2006). Over the business cycle, shocks cause changes in growth rates and in stationary ratios. The shocks to the goods sector productivity and the money supply growth rate are standard, while the third shock to the credit sector productivity exists by virtue of the model’s endogeneity of money velocity via a micro-evidence based (Hancock, 1985) constant-returns-to-scale (CRS) production of exchange credit as in Clark (1984). This credit technology allows for a unique equilibrium between money and credit use even though they are perfect substitutes in exchange for the consumer; by including the deposited funds as an “additional factor” of production this gives a rising marginal cost per unit of deposits that equals the marginal cost of money: the nominal interest rate.¹

The shocks occur at the beginning of the period, observed by the consumer before the decision process, and follow a vector first-order autoregressive process. For goods sector productivity, z_t , the money supply growth rate, u_t , and bank sector productivity, v_t :

$$Z_t = \Phi_Z Z_{t-1} + \varepsilon_{Zt}, \quad (1)$$

where the shocks are $Z_t = [z_t \ u_t \ v_t]'$, the autocorrelation matrix is $\Phi_Z = \text{diag}\{\varphi_z, \varphi_u, \varphi_v\}$ with $\varphi_z, \varphi_u, \varphi_v \in (0, 1)$ as autocorrelation parameters, and the shock innovations are $\varepsilon_{Zt} = [\varepsilon_{zt} \ \varepsilon_{ut} \ \varepsilon_{vt}]' \sim N(\mathbf{0}, \Sigma)$. The general structure of the

¹ Solving this equilibrium problem is discussed as far back as King and Plosser (1984). Both English (1999) and Gillman (2000), for example, have decentralized bank sectors based on a cash constraint over a goods continuum as in Gillman (1993), with related transaction cost technologies. These have unique money/credit equilibria, but not bank production functions as in Clark (1984). Gillman and Kejak (2004) have a related sectoral bank decentralization but leave implicit the role of deposits. Without explicit bank deposits, these papers miss how bank profit equals the interest return to deposits.

second-order moments is assumed to be given by the variance–covariance matrix Σ , with standard deviations of σ_{ε_z} , σ_{ε_u} , and σ_{ε_v} . These shocks affect the economy as described below.

A representative consumer has expected lifetime utility from consumption of goods, c_t , and leisure, x_t ; with $\beta \in (0, 1)$ and $\theta, \Psi > 0$, this is given by

$$U = E_0 \sum_{t=0}^{\infty} \beta^t \frac{(c_t x_t^\Psi)^{1-\theta}}{1-\theta}. \tag{2}$$

Output of goods, y_t , and increases in human capital h_t , are produced with physical capital k_t and effective labor each in Cobb–Douglas fashion; the bank sector produces exchange credit using labor and deposits d_t as inputs. Let s_{Gt} and s_{Ht} denote the fractions of physical capital that the agent uses in the goods production (G) and human capital investment (H), whereby

$$s_{Gt} + s_{Ht} = 1. \tag{3}$$

The agent allocates a time endowment of one amongst leisure, x_t , labor in goods production, l_{Gt} , time spent investing in the stock of human capital, l_{Ht} , and time spent working in the bank sector (Q), denoted by l_{Qt} :

$$l_{Gt} + l_{Ht} + l_{Qt} + x_t = 1. \tag{4}$$

Output of goods can be converted into physical capital, k_t , without cost and so is divided between consumption goods and investment, denoted by i_t , gross of capital depreciation, where the depreciation rate is $\delta_K \in [0, 1]$. Thus, the capital stock used for production in the next period is given by

$$k_{t+1} = (1-\delta_K)k_t + i_t = (1-\delta_K)k_t + y_t - c_t. \tag{5}$$

The human capital investment is CRS-produced using capital $s_{Ht}k_t$ and effective labor $l_{Ht}h_t$:

$$H(s_{Ht}k_t, l_{Ht}h_t) = A_H(s_{Ht}k_t)^{1-\eta}, \text{ with } A_H > 0 \text{ and } \eta \in (0, 1) \tag{6}$$

And the human capital flow constraint, with depreciation rate $\delta_H \in [0, 1]$, is

$$h_{t+1} = (1-\delta_H)h_t + H(s_{Ht}k_t, l_{Ht}h_t). \tag{7}$$

With w_t and r_t denoting the real wage and real interest rate, the consumer receives nominal income of wages and rents, $P_t w_t (l_{Gt} + l_{Qt}) h_t$ and $P_t r_t s_{Gt} k_t$, a nominal transfer from the government, T_t , and dividends from the bank.

The consumer buys shares in the bank by making deposits of income at the bank. Each dollar deposited buys one share at a fixed price of one, and the consumer receives the residual profit of the bank as dividend income in proportion to the number of shares (deposits) owned. Denoting the real quantity of deposits by d_t , and the dividend per unit of deposits as R_{Qt} , the consumer receives a nominal dividend income of $P_t R_{Qt} d_t$. The consumer also pays to the bank a fee for credit services, whereby one unit of credit service is required for each unit of credit that the bank supplies the consumer for use in buying goods. With P_{Qt} denoting the nominal price of each unit of credit, and q_t the real quantity of credit that the consumer uses in exchange, the consumer pays $P_{Qt} q_t$ in credit fees and buys $P_t q_t$ in goods with credit.

With total goods expenditures, $P_t c_t$, physical capital investment, $P_t k_{t+1} - P_t (1-\delta_K)k_t$, and investment in cash for purchases, $M_{t+1} - M_t$, the consumer's budget constraint is

$$P_t w_t (l_{Gt} + l_{Qt}) h_t + P_t r_t s_{Gt} k_t + P_t R_{Qt} d_t + T_t \geq P_{Qt} q_t + P_t c_t + P_t k_{t+1} - P_t (1-\delta_K)k_t + M_{t+1} - M_t. \tag{8}$$

The consumer can purchase the goods by using either money M_t or credit services. With the lump sum transfer of cash T_t coming from the government at the beginning of the period, and with money and credit equally usable to buy goods, the consumer's exchange technology is

$$M_t + T_t + P_t q_t \geq P_t c_t. \tag{9}$$

Since all cash comes out of deposits at the bank, and credit purchases are paid off at the end of the period out of the same deposits, the total deposits are equal to consumption. This gives the constraint that

$$d_t = c_t. \tag{10}$$

The bank produces credit that is available for exchange at the point of purchase. The bank determines the amount of such credit by maximizing its dividend profit subject to the labor and deposit costs of producing the credit. The production of credit uses a constant returns to scale technology with effective labor and deposited funds as inputs (physical capital is omitted as a factor for simplicity). This follows the “financial intermediation approach” (Matthews and Thompson, 2008) that is dominant in the banking literature, which was started by Clark (1984). In particular, with $A_Q > 0$ and $\gamma \in (0, 1)$,

$$q_t = A_Q e^{v_t} (l_{Qt} h_t)^\gamma d_t^{1-\gamma}, \tag{11}$$

where $A_Q e^{v_t}$ is the stochastic factor productivity.²

² This “banking time” model can be interpreted as a special case of the shopping time model: substituting q_t from Eq. (11) into Eq. (9), and for d_t from Eq. (10), and solving for the effective banking time as $l_{Qt} h_t = [(c_t - M_{t+1}/P_t)/(A_Q e^{v_t} c_t^{1-\gamma})]^{1/\gamma}$, then $l_{Qt} h_t = g(M_{t+1}/P_t, c_t)$, with $g_1 < 0$ and $g_2 > 0$, as in a shopping time model. However, there is no Feenstra (1986) equivalence to a standard money-in-the-utility function model in that h_t would enter the utility function, as seen by solving for the raw bank time $l_{Qt} = g(M_{t+1}/P_t, c_t)/h_t$, substituting for l_{Qt} in the allocation of time constraint (4), solving for x_t from this time constraint and substituting into the utility function.

Subject to the production function in Eq. (11), the bank maximizes profit Π_{Qt} with respect to the labor l_{Qt} and deposits d_t :

$$\Pi_{Qt} = P_{Qt}q_t - P_t w_t l_{Qt} h_t - P_t R_{Qt} d_t. \quad (12)$$

Equilibrium implies that

$$\frac{P_{Qt}}{P_t} = \frac{w_t}{\gamma A_Q e^{\nu_t} \left(\frac{l_{Qt} h_t}{d_t} \right)^{\gamma-1}}, \quad (13)$$

$$\frac{P_{Qt}}{P_t} = \frac{R_{Qt}}{(1-\gamma) A_Q e^{\nu_t} \left(\frac{l_{Qt} h_t}{d_t} \right)^{\gamma}}. \quad (14)$$

Eqs. (13) and (14) indicate that the marginal cost of credit P_{Qt}/P_t equals the marginal factor prices divided by the marginal factor products.

The firm maximizes profit Π_t , given by $\Pi_t = y_t - w_t l_{Gt} h_t - r_t s_{Gt} k_t$, subject to a King and Rebelo (1990) technology in effective labor and capital:

$$y_t = A_G e^{z_t} (s_{Gt} k_t)^{1-\alpha} (l_{Gt} h_t)^{\alpha}, \text{ with } A_G > 0 \text{ and } \alpha \in (0, 1). \quad (15)$$

The first order conditions for the firm's problem are

$$w_t = \alpha A_G e^{z_t} \left(\frac{s_{Gt} k_t}{l_{Gt} h_t} \right)^{1-\alpha}, \quad (16)$$

$$r_t = (1-\alpha) A_G e^{z_t} \left(\frac{s_{Gt} k_t}{l_{Gt} h_t} \right)^{-\alpha}. \quad (17)$$

It is assumed that the government policy includes sequences of nominal transfers which satisfy

$$T_t = \Theta_t M_t = (\Theta^* + e^{u_t-1}) M_t, \quad \Theta_t = [M_{t+1} - M_t] / M_t, \quad (18)$$

where Θ_t is the growth rate of money and Θ^* is the stationary growth rate of money.

The equilibrium can be defined by writing the representative agent's optimization problem recursively as

$$\mathcal{V}(s) = \max_{c, x, l_G, l_H, l_Q, s_G, s_H, q, d, k', h', M'} \left\{ \frac{(c x^\psi)^{1-\theta}}{1-\theta} + \beta E \mathcal{V}(s') \right\} \quad (19)$$

subject to the conditions (3)–(10), where the state vector of the economy is denoted by $s = (k, h, M, z, u, v)$ and a prime (') indicates the next-period values. A competitive equilibrium consists of a set of policy functions $c(s)$, $x(s)$, $l_G(s)$, $l_H(s)$, $l_Q(s)$, $s_G(s)$, $s_H(s)$, $q(s)$, $d(s)$, $k'(s)$, $h'(s)$, $M'(s)$, pricing functions $P(s)$, $w(s)$, $r(s)$, $R_Q(s)$, $P_Q(s)$, a transfer function $T(s)$, a value function $\mathcal{V}(s)$, and initial conditions k_0, h_0, M_0 such that

- (i) given $s, P(s), w(s), r(s), R_Q(s), P_Q(s)$, and $T(s)$, the consumer solves the optimization problem in Eq. (19);
- (ii) given $s, P(s), w(s), R_Q(s)$, and $P_Q(s)$, the bank maximizes profit in (12) subject to Eq. (11);
- (iii) given $s, P(s), w(s), r(s)$ the goods producer maximizes profit subject to Eq. (15);
- (iv) the goods, money and credit markets clear, in Eqs. (5), (9), (11), (15) and (18).

2.1. Balanced-growth path equilibrium

As derived from the equilibrium above (see A.4 for the first-order conditions), a partial set of equilibrium conditions along the balanced-growth path (BGP) are given here to describe the deterministic balanced-growth path equilibrium, and how inflation affects it. The balanced-growth rate is denoted by g , with c_t, k_t, h_t, q_t, d_t , and M_{t+1}/P_t all growing at the rate g ; nominal prices grow at the stationary inflation rate denoted by π . Dropping time subscripts on stationary variables, the BGP conditions include

$$1 + \frac{P_{Qt}}{P_t} = 1 + R = (1 + \pi)(1 + r - \delta_K), \quad (20)$$

$$\frac{x}{\alpha c_t} = \frac{1 + \dot{R}}{w h_t}, \quad (21)$$

$$\dot{R} = \left(1 - \frac{q_t}{d_t} \right) R + \left(\frac{q_t}{d_t} \right) \gamma R, \quad (22)$$

$$\frac{q_t}{d_t} = (A_Q)^{1/(1-\gamma)} \left(\frac{\gamma R}{w} \right)^{\gamma/(1-\gamma)}, \quad (23)$$

$$r_H \equiv \eta A_H \left(\frac{s_H k_t}{l_H h_t} \right)^{(1-\eta)} (1-x), \quad (24)$$

$$(1+g)^0 = \beta(1+r_H-\delta_H) = \beta(1+r-\delta_K). \quad (25)$$

The solution for consumption-normalized money demand (inverse velocity), $1-q_t/c_t$, is derived from Eqs. (9), (10) and (23); the consumption velocity of money, denoted by V_{c_t} , is given by

$$V_{c_t} \equiv \frac{c_t}{\frac{M_{t+1}}{P_t}} = \frac{1}{1-\frac{q_t}{c_t}}. \quad (26)$$

This rises at an increasing rate as the nominal interest rate rises. In particular, the interest elasticity of normalized money, $(M_{t+1}/P_t)/c_t$, is equal to $-(q_t/c_t)\gamma/[(1-q_t/c_t)(1-\gamma)]$, and this rises in magnitude as the interest rate increases and causes more credit use. The relative price of credit P_{Q_t}/P_t equals its marginal cost and by Eq. (20) this equals the nominal interest rate R . At the first best optimum, R equals zero and no credit is used. As inflation rises, the agent substitutes from goods towards leisure while equalizing the margin of the ratio of the shadow price of goods to leisure, $x/(\alpha c_t) = [1+\bar{R}]/(wh_t)$, in Eq. (21). Here \bar{R} , as given in Eq. (22), is the average exchange cost per unit of output; this equals the average cost of using cash, R , weighted by $1-q_t/d_t$ and the average cost of using credit, γR , weighted by q_t/d_t . Eq. (23) gives the solution for q_t/d_t , and in turn velocity (Eq. (26)), from Eqs. (11), (13), and (20) (the dividend rate R_{Q_t} follows from Eq. (14)). Eqs. (24) and (25) indicate the subsequent growth effects that are important in identifying the money and credit shocks within the endogenous growth framework. Inflation-induced increases in leisure decrease the human capital return of Eq. (24), and lower the growth rate in Eq. (25). The use of more credit to avoid inflation instead of using leisure means that leisure increases by less and the growth rate falls by less (Gillman and Kejak, 2005). Therefore a shock to the money supply causes higher inflation, more credit use, less money demand, higher velocity and less growth. If such higher inflation and increased credit use are correlated with productivity innovation in the credit sector, then this leads to a high correlation between the money and credit productivity shocks as identified by the model.

2.2. Exogenous growth

For comparison, a way of introducing exogenous growth is by letting human capital follow the exogenous trend given by g : $h_{t+1} = h_t(1+g)$, and with $l_{Ht} = s_{Ht} = 0$. As in endogenous growth, the sectoral productivity parameters A_G and A_Q are constant and also both the income and consumption velocities of money are constant along the BGP, while all growing variables grow at the same rate of g .

3. Model simulation

By normalizing the variables that grow along the deterministic steady state, and then log-linearizing the all model equilibrium conditions around this normalized deterministic steady state, we get a stochastic linear system of equations. Here we normalize by dividing by the human capital stock h_t . This system of equations is in terms of k_t/h_t and the three shocks; it is solved by using standard techniques as in Hartley et al. (1997).

3.1. Calibration

Table 1 presents the parameters for the calibration which are chosen in order to match the Table 2 target values that are average annual values from US time series for 1919–2004 and that reflect issues raised by Gomme and Rupert (2007), in their two-sector RBC model; our human capital sector is a related second sector. The capital share in the goods sector is set at $1-\alpha = 0.36$ as in Jones et al. (2005), the annual discount factor is set at $\beta = 0.96$, and log-utility is assumed so that $\theta = 1$. The US average annual output growth rate g is set at 2.4% as in the data. The baseline investment to output ratio target value is $i/y = 0.26$, implying the annual depreciation rate $\delta_K = 0.031$ and the real interest rate net of depreciation of $r-\delta_K = 0.067$.³ The rate of depreciation of human capital is set at $\delta_H = 0.025$, as in Jones et al. (2005) and Jorgenson and Fraumeni (1989). The allocation of time is similar to Gomme and Rupert (2007), with the working time set at $l_G = 0.2482$ and leisure at $x = 0.55$. Time in human capital investment is set at $l_H = 0.2$. Given l_H , g , and δ_H implies the capital to effective labor ratio in the human capital sector, its capital share of $\eta = 0.83$, and in turn an $A_H = 0.21$ and a leisure utility weight of $\Psi = 1.84$.

In the banking sector we set the value of the inverse of the consumption velocity of M1 money, $M/(Pc)$, equal to the average annual value for the period 1919–2004, which is 0.38. The average annual inflation rate, π , over the same period is 2.6% which implies that the annual money growth Θ^* is equal to 5%. Using an approximate cost of an exchange credit card (American Express) at \$100, and the per capita annual consumption expenditure, $c = \$15,780$, both at 2006 prices, the

³ For comparison $i/y = 0.13$ in Gomme and Rupert (2007) for postwar market structures, equipment and software; including their consumer durables increases this by 0.10, and housing adds another 0.056, for a total of 0.286.

Table 1
Parameters of calibration.

<i>Preferences</i>		
θ	1	Relative risk aversion parameter
Ψ	1.84	Leisure weight
β	0.96	Discount factor
<i>Goods production</i>		
α	0.64	Labor share in goods production
δ_K	0.031	Depreciation rate of goods sector
A_G	1	Goods productivity parameter
<i>Human capital production</i>		
η	0.83	Labor share in human capital production
δ_H	0.025	Depreciation rate of human capital sector
A_H	0.21	Human capital productivity parameter
<i>Banking sector</i>		
γ	0.11	Labor share in credit production
A_Q	1.1	Banking productivity parameter
<i>Government</i>		
Θ^*	0.05	Money growth rate
<i>Shocks processes</i>		
<i>Autocorrelation parameters</i>		
φ_{ε_z}	0.84	Production productivity
φ_u	0.74	Money growth rate
φ_v	0.73	Banking productivity
<i>Standard deviation of shock innovations</i>		
σ_{ε_z}	0.77	Production productivity
σ_{ε_u}	0.50	Money growth rate
σ_{ε_v}	1.16	Banking productivity

Table 2
Target values of calibration.

g	0.024	Avg. annual output growth rate
π	0.026	Avg. annual inflation rate
l_G	0.2482	Labor used in goods sector
l_H	0.20	Labor used in human capital sector
l_Q	0.0018	Labor used in banking sector
i/y	0.26	Investment–output ratio in goods sector
$M/(Pc)$	0.38	Normalized money

share of the labor in the banking sector is $\gamma = 100/[Rc(1-[M/(Pc)])] = 0.11$. A similar calibration of this labor share is made in Benk et al. (2008). It implies that only 11% of the interest return R is used up in the process of producing the exchange credit, which is a result focused on in Gillman and Kejak (2009). And it follows that the marginal cost per unit of credit is an upward sloping, and for any $\gamma < 0.5$, convex curve; so with this calibration the supply curve has its typical shape of a marginal cost rising at an increasing rate.

Table 1 also includes the parameters characterizing the shock processes of Eq. (1); these are chosen through an iterative process by which the assumed shock parameters converge with the actual shock parameters that are in turn estimated from the constructed shock processes described in A.1. In particular, estimated parameters are inputted back into the model, shocks are re-constructed and parameters re-estimated until convergence is achieved in the parameter structure. For comparison, the exogenous growth version of the model without human capital investment assumes the same parameters as those used for the baseline endogenous growth model, except that the target labor time l_G increases to include the targeted time l_H in human capital investment.

3.2. Effect of shocks on velocity

Fig. 2 illustrates the impulse responses of income velocity (vel in figure) when faced with a temporary 1% increase in the credit shock (CR), goods productivity shock (PR), and money shock (M). All three shock cause velocity to rise initially, with a gradual decrease back to equilibrium for the credit and money shocks, and some decrease in velocity from the PR shock after the initial increase. The productivity shock increases velocity mainly by increasing temporary output; velocity falls after a while, before returning to the original equilibrium, as the increased goods productivity decreases the price of goods

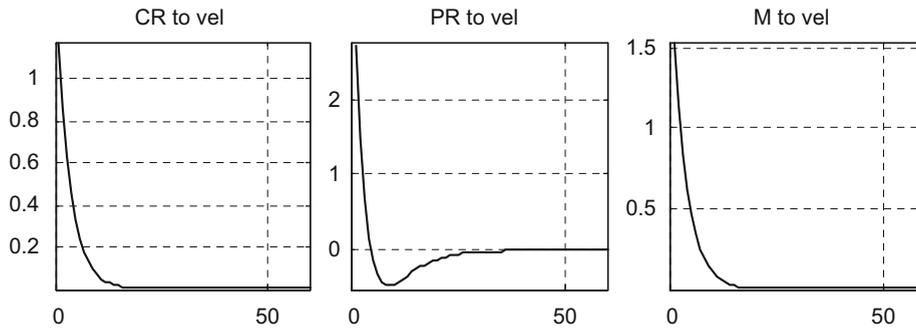


Fig. 2. Impulse responses of income velocity of money.

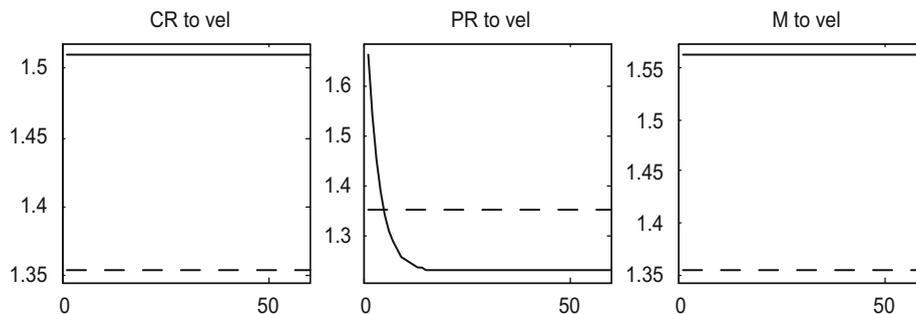


Fig. 3. Transition dynamics of income velocity of money.

relative to labor, so that inflation decreases and money demand increases. The money shock causes a jump in the price level, inflation, and interest rates and decreases real money demand; the credit shock makes the marginal cost of credit lower, inducing a decrease in money demand and an increase in credit use.

Consider in contrast when there is a shock that does not dissipate, so that the shock is permanent, and then velocity transitions from one BGP to another one. Assuming a 10% permanent increase in each of the three shock variables, Fig. 3 shows the original (dashed line) BGP equilibrium and the movement to the new BGP equilibrium (solid line). The figure shows the goods sector shocks causes income velocity to initially increase and then fall down to a new lower steady state, while the money and credit shocks cause velocity to rise with virtually no transition dynamics. The goods sector shock causes income to rise initially while prices are gradually decreased; with less inflation, more money is used and income velocity eventually falls. The money shock simply makes inflation higher and increases velocity, while the credit shock makes credit less expensive and so decreases money demand and increases velocity.

3.3. Simulations

Table 3 presents US data (see A.3) stylized facts and model simulations, in terms of moments of a set of variables for the period 1919–2004, where the data series have been detrended using the Christiano and Fitzgerald (2003) asymmetric frequency filter with a band of 2–86 years (where 86 is the sample size). And in the simulations, the consumption and investment level variables are normalized by human capital. The table shows that simulated relative volatilities of consumption and investment ratios are 0.51 and 2.97, compared to data of 0.64 and 4.09. Output growth volatility is 0.29 compared to 0.48 for GDP data. And simulated consumption and investment correlation with output is 0.71 and 0.94 versus 0.56 and 0.53 in the data.

On the monetary side, the simulated volatility of the income velocity of money is 1.21 as compared to 1.39 in the data, which means the model explains 87% of the volatility found in the GDP velocity of $M1$ that exists in the data. The normalized real money volatility is 1.60, almost identical to 1.61 in the data. The income velocity's correlation with output is -0.03 , compared to the data's 0.06 . Real money correlation with output is 0.65 compared to the data value of 0.47 ; and with output growth it is 0.26 versus 0.14 . Not shown, the simulated correlation of inflation with money growth is 0.49 , as compared to 0.42 in the data.

4. Decomposing the effect of shocks on velocity

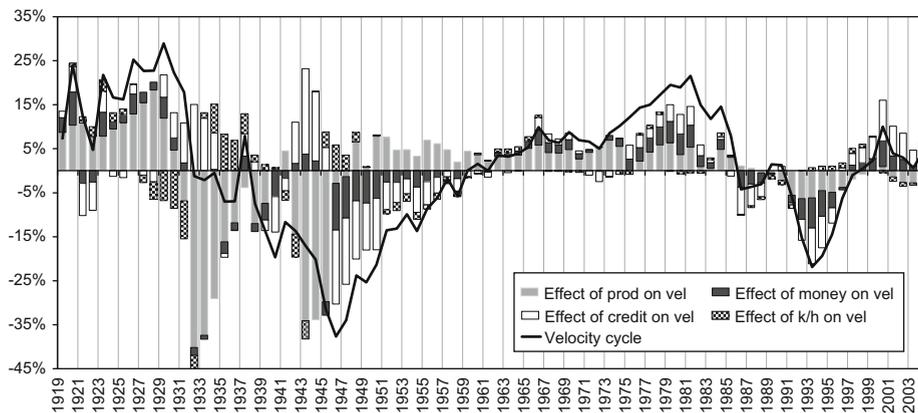
We construct the shocks by adapting the methodology of Ingram et al. (1994) for this endogenous growth model, a procedure described in Benk et al. (2008). Here we solve the equilibrium solution for each control variable as a function of

Table 3

US business cycle facts, 1919–2004, and simulations.

Simulation results data: 1919–2004	Relative volatility		Output correlation		Output growth rate correlation	
	Simulated	Data	Simulated	Data	Simulated	Data
Consumption	0.51	0.64	0.71	0.56	0.30	0.01
Investment	2.97	4.09	0.94	0.53	0.31	0.17
Output growth rate	0.29	0.48	0.35	0.20	1.00	1.00
Employment	0.73	0.75	0.84	0.91	0.26	0.30
Income velocity of money	1.21	1.39	−0.03	0.06	−0.05	0.01
Normalized real money	1.60	1.61	0.65	0.47	0.26	0.14

Note: See appendix for data sources. All data series represent the cyclical component of the data filtered with the Christiano and Fitzgerald (2003) asymmetric frequency filter with a band of 2–86 years (86 = sample size). Series are in logs except those that represent rates. Relative volatility is measured as the ratio of standard deviation of the series to the standard deviation of GDP.

**Fig. 4.** Effect of shocks and k/h on US income velocity; endogenous growth model, 1919–2004.

the endogenous state variable and the three shocks, gather time series for a set of choice variables and use this data to solve for each shock at each time t ; see A.1. With three such time series the shocks are exactly identified. We use more than three series to base the shocks on more information, thereby over-identifying the shocks, and then estimate the shocks using a minimum distance approach, described in Benk et al. (2008). A difference from previous work is that here we use annual data and a band pass filter that takes out only the 86 year trend, so as to leave in longer run cycles along with business cycles.

4.1. Effect on velocity levels

The effects of these shocks on the income velocity of money can be devised by decomposing the cyclical component of velocity into the contributions of productivity, money and credit shocks. A linear decomposition can be done by using the solution of the model that writes every model variable as a linear function of the state vector $s = (k_t/h_t, z, u, v)$; where for any ξ , $\hat{\xi} \equiv \log(\xi) - \log(\bar{\xi})$ is the percentage deviation of ξ from its steady state value $\bar{\xi}$. For the income velocity of money, denoted by V_t , it follows that

$$\hat{V}_t \equiv \frac{\widehat{y_t P_t}}{\widehat{M_{t+1}}} = \phi_k k_t/h_t + \phi_z z_t + \phi_u u_t + \phi_v v_t + e_t^V, \quad (27)$$

where $\phi_z z_t$, $\phi_u u_t$ and $\phi_v v_t$ indicate the contribution of productivity, money and credit shocks to the cyclical component of velocity.⁴ Since we use more variables than shocks, the model does not perfectly fit the data series used to construct the shocks, leaving an error term e_t^V .

Velocity decompositions based on Eq. (27) are shown in Fig. 4, along with the Christiano and Fitzgerald (2003) 86 year detrended velocity itself. Much of the movement of the velocity cycles over time is captured by the shocks of the model, with less success during the 1930s depression and more success since the end of WWII. For example, Fig. 4 shows that

⁴ For k/h we are not aware of any human capital estimate back to 1919, and so assume a smooth trend for h that has no effect on the filtered data.

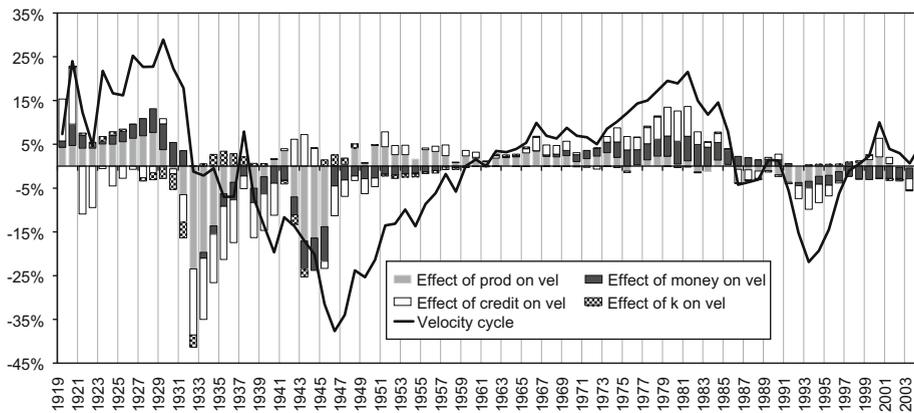


Fig. 5. Effect of shocks and k/h on US income velocity; exogenous growth model, 1919–2004.

productivity shocks contributed to a significant amount of the velocity changes after WWII, but that the money and credit shocks were also important such as during the 1970s inflation and the post-2001 velocity movement.

The effect of shocks on the level of filtered income velocity can also be computed for the exogenous growth version of the model (Fig. 5). Comparison of Figs. 4 and 5 gives a distinct sense in which the endogenous growth model is able to explain more of the actual filtered velocity level than the exogenous growth model. This appears true in the 1920s, from 1939–1959, 1961–1970, and since 1990. The actual shocks of the endogenous growth versus exogenous growth models are compared in Fig. 7 in A.1; endogenous growth forces more correlation between the money and credit shocks. There is more “smoothing” for the money shock with exogenous growth; and the credit shock behavior during the 1930s differs between the two models.

4.2. Effect on velocity volatility

We decompose the fluctuations in the cyclical component of the GDP velocity of M1 money (which is from the data) and show how much of the variance is explained within each subperiod by each of the model’s three shocks, the productivity (PR), money (M) and credit (CR) shocks, and across business cycle (BC) and long run (LR) frequencies such as in Levy and Dezhbakhsh (2003) (see A.2). With a variation on Ingram et al. (1994), we take an unweighted average over all six possible orderings of the three shocks. Table 4 reports variance decompositions for the entire 1919–2004 period, for 1919–1935 (Roaring 20s-depression), for 1936–1954 (recovery-WWII), for 1955–1982 (postwar and high inflation), and for 1983–2004 (Moderation). Here the variance-covariance matrices have been estimated separately for each subperiod so as to obtain simulated series and decompositions that differ by subperiods.

Table 4 shows the US (M1) income velocity variance decomposition results. The columns show the fraction of the data variance, by frequency, that is explained by each shock, for both endogenous and exogenous growth models. For example, with endogenous growth, in the subperiod of 1919–1935, the model explains a total of $15 + 14 + 13 = 42\%$ of the actual variance found within the BC frequency, and $33 + 34 + 37 = 104\%$ of the variance within the LR frequency, for a total of 146% of the variance. Therefore the 1919–1935 total model volatility is 46% more than in the data. Similarly the model explains 263% of the volatility for 1936–1954, 109% for 1955–1982, 54% for 1983–2004, and 76% for the whole period, 1919–2004.

The standard deviation and correlation of the shocks differ somewhat between the endogenous and exogenous growth versions, for the whole period, and when divided by subperiod (details not reported). Both models have a standard deviation of the productivity shock that drops significantly post-1954. In both models, the standard deviation of the money shock is rather stable across all four subperiods, while the standard deviation of the credit shock is lower in the second two subperiods than the first two, as with the goods productivity shock. Focusing on endogenous growth, in the whole 1919–2004 period, the standard deviations are 0.48, 0.28, and 0.68 for PR, M, and CR. And the credit shock standard deviation doubles from 0.21 in 1955–1982 to 0.40 in 1983–2004. In the whole period, the money and credit shocks have a 0.75 correlation, while credit and goods productivity shocks have a -0.51 correlation, and money and goods productivity have a 0.17 correlation. However, there is a negative correlation of credit with goods productivity in each of the first three subperiods but a positive one during 1983–2004. And money and credit shocks have a high positive correlation (above 0.90) in the latter three subperiods but 0.18 in 1919–1935.

4.3. Discussion of results

The results show the importance of the contribution of the shocks within the long run frequency as well as the importance individually of each of the three shocks. A positive goods productivity shock works mainly on the income term

Table 4

Decomposition of variance of velocity by frequency, 1919–2004.

US (M1) income velocity variance decomposition	Endogenous			Exogenous		
	PR (%)	M (%)	CR (%)	PR (%)	M (%)	CR (%)
1919–1935						
BC	15	14	13	9	7	7
LR	33	34	37	23	23	75
1936–1954						
BC	1	19	18	1	30	3
LR	59	77	89	37	6	25
1955–1982						
BC	0	13	14	0	15	2
LR	4	33	45	4	24	69
1983–2004						
BC	2	11	12	2	10	8
LR	25	2	2	17	0	27
1919–2004						
BC	1	16	15	1	19	6
LR	13	14	17	7	1	42

in the income velocity of money, by increasing the economy's temporary output/income. This causes output to rise relative to consumption since consumption follows the permanent income and only increases somewhat when temporary income rises. Therefore the income to consumption ratio rises.

With a standard exchange constraint, and only cash used to buy goods, then the story of income velocity ends with the income to consumption ratio since money demand always equals consumption demand. With credit available, the consumer can substitute away from money towards credit. The money and credit shocks affect this substitution and thereby affect mainly the money to consumption ratio. A positive money shock raises inflation and causes substitution towards credit; this shock also decreases the growth rate of income because of the inflation tax effect. A positive credit productivity shocks works similarly in that substitution occurs from money to credit, but in reverse has a positive effect on income and its growth rate. The main effect on velocity from the money and credit shocks is the substitution effect rather than the income effect, and both affect velocity similarly in magnitude (as in Figs. 2 and 3).

When the productivity shock explains a greater amount of velocity volatility than the money and credit shocks, it means that the period is characterized more by simple changes in temporary income that little affect the consumption demand. This type of velocity volatility from goods productivity shocks is more of the “normal” type associated with the RBC real economy model. However, when there are significant variations in the money supply, such as in response to big debt increases from war or bank crises, then there can be more of the traditional monetary type of velocity volatility. And these money shocks either stimulate credit use or credit may be suppressed and cannot respond. In the former case, the velocity rises by more because of both less money and more credit use. If credit is suppressed, such as during the depression when banks failed, and in the recent bank crisis, then velocity volatility would be expected to be explained somewhat more by money than by credit shocks.

4.3.1. US velocity volatility

Results show for endogenous growth that money and credit shocks contribute similar amounts to velocity volatility. In the first two subperiods before 1955, money contributed marginally more at the business cycle frequency than credit, while in the last two subperiods credit contributed slightly more than money to volatility. This is not inconsistent with credit being more constrained in the first two subperiods and less constrained in the second two, when there was steady financial innovation as international capital markets expanded after WWII and as deregulation took place in the 1970s and 1980s. The large role of money and credit shocks in the LR frequency indicate that these shocks added up to have an effect of greater permanence than found in the typical business cycle frequency. This can be interpreted as the long run inflation tax nature of money shocks being more important than the business cycle effects, all the way from 1919 to 1982. During the 1983–2004 “Moderation” and inflation targeting policy, the long term effects of credit and money shocks were not important. Instead the goods productivity effects dominated the long term effects on velocity, in reverse of 1955–1982.

4.3.2. US velocity levels

The effect of shocks on the magnitude of filtered velocity levels shows an additional dimension to the volatility effects. In Fig. 4, for example in the 1990s, the level effects are nicely explained by the shocks: during high growth and low inflation, velocity levels moved down as explained by goods productivity lowering inflation and by money supply policy lowering inflation. High credit shock components of velocity at the end of the 1990s and up to 2004 indicate that credit is starting to have a large level effect, while the contribution of goods productivity is receding, a possible warning sign of the ending of the stable period.

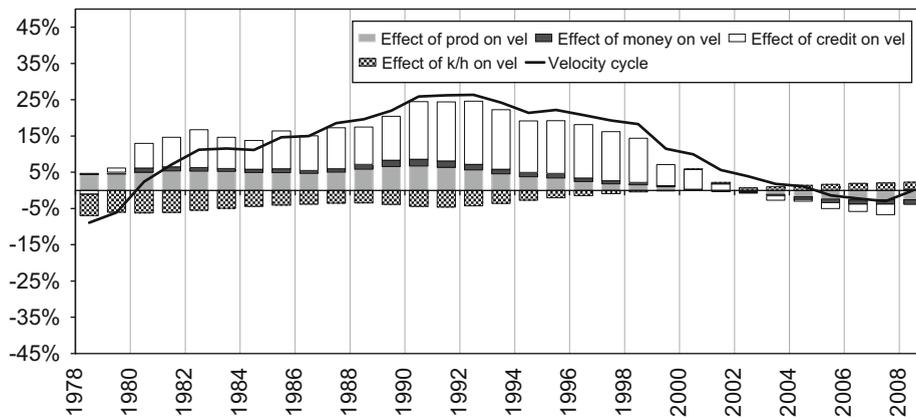


Fig. 6. Effect of shocks on UK income velocity; endogenous growth model, 1978–2008.

Changes in the state variable, k/h (indicated by the dotted portions on the graph) affect the level of filtered velocity negligibly post-1955 but significantly pre-1955. The transition dynamics indicate that state variable changes occur mainly from changes in the goods productivity shock, and only slightly from the money and credit shocks (details not shown). This suggests that the high contribution of the state variable to pre-1955 velocity levels is mainly from erratic goods productivity shocks. The lesser effect of the state variable on velocity after 1955 in turn reflects how the goods productivity shock is more stable; indeed the standard deviation of the goods productivity shock falls substantially in the two post-1955 subperiods to 0.07 and 0.11, chronologically, from the two pre-1955 subperiod values of 0.74 and 0.68.

5. Extension to the UK

The analysis can be extended to the UK for a more limited time period. The UK data for the financial sector exists starting in 1978, from the Office for National Statistics, and this determines our period of 1978–2008. Rather than ending at the 2004 moderation end-point, this sample includes moving into the recent bank crises.

The income velocity of the UK $M0$ aggregate increases on average by 2.1% per year over the period, compared to 1.25% average per year increase for the US $M1$ income velocity.⁵ The calibration of the model is kept similar to the US, except with the money supply growth rate at 0.075, inflation at 0.05, and normalized money at 0.33. The shock processes for this more limited period have higher autocorrelations than for the US, and lower standard deviations.⁶ The same 2–86 filter on data is used for more direct comparison, although a 2–31 filter could also be used and we expect the results to be similar.

Given the greater persistence of the shocks for the UK, the impulse responses for the UK (not shown) reflect a slower return to the steady state. The actual computed shocks (not shown) are positively correlated at 0.89 between money and credit shocks, 0.94 between money and goods productivity shocks, and 0.73 between credit and goods productivity shocks. The goods productivity shock goes from a positive value to becoming steadily negative after 2001 for the endogenous growth model. The credit productivity shock also turns slightly negative around that time, while the money shock is more correlated with the goods productivity shock.

Fig. 6 for the endogenous growth model shows that the model seems to capture the level of the filtered UK $M0$ income velocity well, while the exogenous growth model does not appear to explain the velocity as well (not shown), just as with the US results. Table 5 shows that most of the volatility of the income velocity is explained by the productivity shock in the LR part of the spectrum, and that money plays the next biggest role in explaining volatility.

Comparison of the UK and US endogenous growth results can best be made for the similar periods of 1983–2004 for the US and 1978–2008 for the UK. The productivity shock in both countries during this time plays the biggest role in explaining velocity volatility, while the credit shock contributes more to the volatility in the US than in the UK. For both the US and UK the goods and credit productivity shocks show up significantly in affecting the level of velocity, in Figs. 4 and 6. This suggests that the credit shock has been important for both US and UK but more stable in the UK. The UK–US results, for the similar time periods, support the interpretation that “normal” goods productivity effects explain most of the volatility, rather than the money and credit effects as in earlier periods in the US. But at the same time, the drop off of the positive credit shocks and the downturn in filtered velocity, from 2000–2003 for the US and for 2001–2007 for the UK, indicate

⁵ $M0$ is used as the money aggregate in the data used to construct the shocks for the UK, and the cyclical component of $M0$ income velocity of money is graphed in Fig. 6. The UK $M1$ aggregate includes overnight deposits of banks that are much higher proportionately than in the US or the Euro area, making $M1$ not so comparable to the broader US monetary aggregates. With the UK money to consumption ratios of the aggregates for the period 1978–2008 being $M0/(Pc) = 0.06$, $M1/(Pc) = 0.78$, $M2/(Pc) = 1.27$, and $M4/(Pc) = 1.24$, we use an intermediate value between $M0$ and $M1$ of $M/(Pc) = 0.33$ for the UK calibration, which is close to the US value of $M/(Pc) = 0.38$.

⁶ These are $(\rho_z, \rho_u, \rho_v) = (0.96, 0.95, 0.94)$, and $(\sigma_z, \sigma_u, \sigma_v) = (0.022, 0.019, 0.121)$.

Table 5

Decomposition of variance of velocity by frequency, UK.

UK (M0) income velocity variance decomposition	Endogenous			Exogenous		
	PR (%)	M (%)	CR (%)	PR (%)	M (%)	CR (%)
1978–2008						
BC	9	7	2	5	4	0
LR	88	5	1	80	14	8

potential economy-wide fragility. With hindsight it can be postulated that these credit shocks as seen through velocity were precursors to the ensuing 2007–2009 credit crisis: consider that related credit shock and filtered velocity patterns emerge before the troughs of the US 1930s depression and the US 1987 recession. In fact, US filtered velocity turned negative in the 1930s, again in 1986–1987, and approached this in 2003; and this happens in the UK in 2004.

6. Comparison to literature

Ingram et al. (1994) show that the variance decomposition of the goods productivity shock is not unique but rather depends upon the ordering of the shocks. We verify that result by considering all possible shock orderings, and since there is no obvious rationale for ordering one shock ahead of another, an average is taken across all orderings. Our work is relatively novel in that the shocks considered are the “standard” monetary RBC shocks of goods productivity and money, plus a more novel shock to the credit sector productivity. This directly extends the Cooley and Hansen (1989) work to when velocity is endogenous, while using this exciting shock construction technique of Ingram et al., in which the world as we know it is the model, and given this, these are the shocks of this world. Plausibility requires that the exchange credit shock is based in a realistically endogenous velocity. The production of exchange credit using the empirically robust Clark (1984) production function provides such a plausible endogenous velocity instead of using the standard cash-only consumption velocity of unity; this makes it a reasonable basis for the credit shock that affects velocity. Although such a credit shock is novel to our own work, a related approach is used in Nolan and Thoenissen (2009). They also have standard goods productivity and money supply shocks, and a third, novel, credit shock. Their alternative credit shock is based in a financial accelerator model, and so is more of an intertemporal credit shock rather than the exchange credit shock of this paper. But as in Ingram et al. (1994) and this paper, they also back out the shocks using the same methodology. They show how their credit shock is an interesting indicator of economic downturns.

Our application here is the first to examine long historical velocity data with this methodology, using a plausible credit shock, and its goal is to show that analysis of the identified shocks offers a way to better understand not just velocity per se but also how velocity foreshadows economic crises, especially bank related crises. It is staking out the case that velocity matters not only as another aspect of the long cyclic experience that a good model should be able to explain in conjunction with other “real” aspects, but also as a possible harbinger of crisis that consequently warrants further study.

Finally, our filtering methodology of including the long run frequency in the data as well as the business cycle frequency is novel in this context and indicates how models that endogenize growth over time may be the best route towards understanding the effect of both the long lasting shocks that indeed should affect growth as well as those shocks that are more defined at business cycle frequencies during which exogenous growth assumptions may be more innocuous.

7. Conclusion

The paper explains US velocity cycles around its 1.25% trend in terms of historically constructed money, credit and goods productivity shocks under endogenous growth. The results show how a significant proportion of the volatility of (86-year band-pass filtered) velocity can be explained with these shocks, in both business cycle and long run frequencies. Applying the model also to the UK for 1978–2008 supports features seen during the US in the 1983–2004 moderation period, such as stable economic periods coinciding with velocity volatility being explained mainly by the goods productivity shock. An interpretation is that as money supply policy fluctuations are stabilized, such as with inflation targeting at low levels of the inflation rate, there is less of a need to use variations in credit to avoid variations in inflation, leaving only normal fluctuations in temporary income from the goods productivity shock to affect velocity volatility. During more unstable periods of monetary policy, the money and credit shocks can swamp the effect of the goods productivity shocks and explain more of the velocity variation. Given the correlation amongst the identified shocks, it is also plausible that more variable money and credit shocks themselves lead to more variable goods productivity shocks.

Understanding velocity volatility provides another dimension, or reflection of activity if you will, that helps explain how we get more stable periods of aggregate activity, such as after 1982. Conjecturing from here towards recent policy experience, this suggests that the US sustained low nominal interest rate policy of 2002–2004, and of 2008 onwards, may be cases described by Friedman (1968) of trying to peg low real interest rates that lead to new eras of unwanted fluctuations in government debt, money supply, and private credit. When huge debt build-ups occur, such as during the

current bank crisis, then the dominance of goods productivity shocks will likely recede as money and credit shocks again take over in importance in explaining velocity. And should this occur, as it seems it is now likely, it portends the advent of greater inflation and even output volatility.

For monetary policy, a “state-dependent” money supply rule that offsets velocity changes in order to target inflation, as in McCallum (1990) and Keynes (1923), could be directed at the band-pass filtered velocity. Such a policy in the US historical sample here would have raised money supply growth substantially during the 1930s depression, the 87 stock crash, the 1991 recession and at the end of the sample in 2004 and beyond, in that filtered velocity has continued to fall. Unlike Taylor interest rate rules with issues of a zero nominal bound, this policy rationalizes the current US policy of “quantitative easing” that involves dramatically increasing the money supply growth rate.

Acknowledgments

The views expressed here are solely those of the authors and do not necessarily reflect the official views of the Fiscal Council of the Republic of Hungary. We are grateful to the referees for excellent critiques and to Jim Bullard for useful comments.

Appendix A

A.1. Shock construction

In this approach we used six variables to construct the economy’s three shocks, shown in Fig. 7, as opposed to five variables in Benk et al. (2008), and with the additional variable being the inflation rate π (others are $M/(Py)$, c/y , i/y , l_Q and the wage rate in banking as a proxy for the marginal product of labor in banking). Using data series on these six variables and computed series on the endogenous state variable, k/h , the three shocks are then identified using a minimum distance approach as in Benk et al. (2008). For robustness, the shock construction is repeated for six variables taken five at a time, four at a time, and three at a time (exact identification). Nearly the same shock series results for all combinations that include π , $M/(Py)$, and either c/y or i/y , and either l_Q (banking hours) or the banking wage rate; this indicates robust shock construction as long as the variables included correspond to the model’s output and banking sectors in which productivity shocks occur, plus inflation which primarily reflects the money shock, plus velocity.

A.2. Variance decomposition by spectral frequency

The variance of velocity is decomposed along a third dimension and it is shown the amount of variance that takes place at business cycle and long run frequencies. The business cycle (BC) frequency band corresponds to cycles of 2–10 years and the long-run (LR) band to cycles of 10 years and longer.

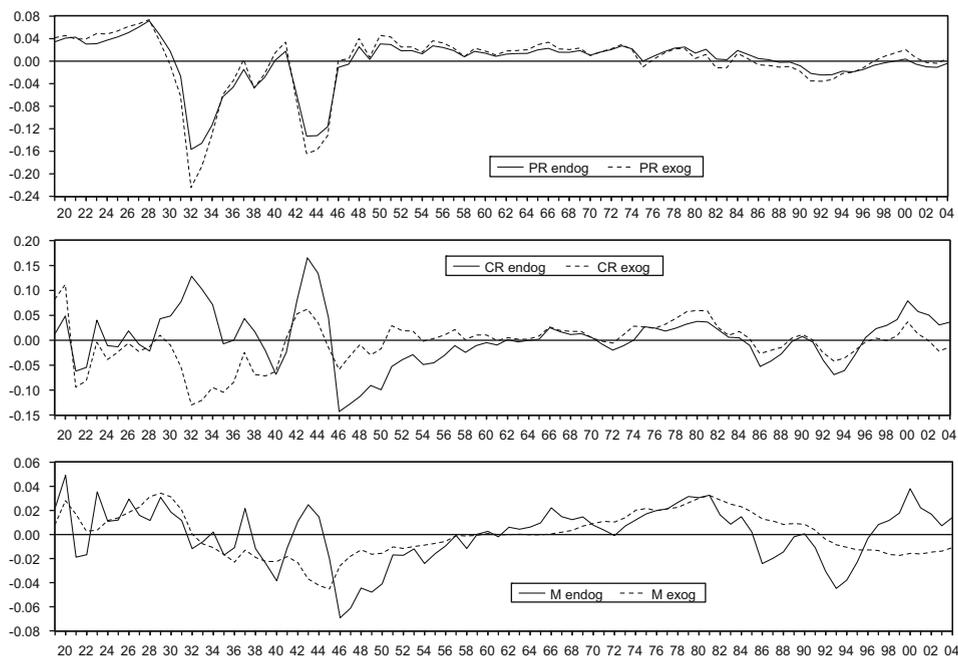


Fig. 7. Productivity (PR), credit (CR) and money (M) shocks.

The proportion of variance of a series due to BC and LR components can be obtained as in Levy and Dezhbakhsh (2003): it amounts to estimating the spectral density of the series, normalizing it by the series variance, and then computing its integral over the corresponding frequency band. If we denote by $f(\omega)$ the spectral density of the series and by σ^2 its variance, then the fraction of variance due to each frequency component is given by $H^{BC} = \int_{2\pi/10}^{2\pi/2} f(\omega)/\sigma^2 d\omega$, $H^{LR} = \int_{2\pi/\infty}^{2\pi/10} f(\omega)/\sigma^2 d\omega$. The frequency bands are determined by the mapping $\omega = 2\pi/p$, where p measures the cycle length (2 or 10 years). Using an alternative, equivalent measure for the fractions of variance (suggested also by Levy and Dezhbakhsh, 2003), this consists of passing the series through a Christiano and Fitzgerald (2003) asymmetric band-pass filter with 2–10 and > 10 year windows, estimating the variance of the filtered series and relating it to the variance of the original series. This procedure is applied to the simulated series of velocity, where simulations have been run by feeding back the estimated variance–covariance structure of the shocks into the model. To assess the fraction of variance explained by each shock in turn at each frequency, we decompose each of the frequency component further, by shocks. The procedure here is similar to Ingram et al. (1994), but requires pre-filtering both the target velocity series and the three shock series to extract the adequate frequency component.

A.3. Data sources

Data used in the paper have been constructed on annual frequency, for the 1919–2004 time period. The main data sources were the Bureau of Economic Analysis (BEA) and the IMF International Financial Statistics (IFS). Series have been extended backwards until 1919 based on the series published in Kuznets (1941), Friedman and Schwartz (1963) (F&S) and the online NBER Macroeconomic History Database (<http://www.nber.org/databases/macrohistory/contents/>) (NBER). The data series are as follows: Gross Domestic Product (BEA, Kuznets); Consumer Price Index (BEA, F&S); Price Index for Gross Domestic Product (BEA, Kuznets); Personal Consumption expenditures (BEA, Kuznets); Gross private domestic investment (BEA, Kuznets); Wage and salary accruals (BEA, Kuznets); Wage and salary accruals, Finance, insurance, and real estate (BEA, Kuznets); Full-time equivalent employees (BEA, Kuznets); Full-time equivalent employees, Finance, insurance, and real estate (BEA, Kuznets); M0 (IFS, NBER); M1 (IFS, NBER); M2 (IFS, NBER); Treasury Bill rate (IFS, NBER).

For the UK the data set has been constructed on annual frequency, for the 1978–2008 period. The length of the sample was limited by the availability of financial sector data that we collected only starting from 1978. The main data sources were the IMF International Financial Statistics (IFS), the UK Office for National Statistics (ONS) and the Bank of England (BoE). The data series are as follows: Gross Domestic Product (IFS); Consumer Price Index (IFS); Household Consumption expenditures (IFS); Gross fixed capital formation (IFS); Labour in banking sector: proxied by the ratio of Jobs in financial intermediation (industry J) (ONS) to Total workforce jobs (ONS); Marginal product of labor in banking: proxied by labour productivity in banking sector, calculated as the ratio of Financial intermediation output (ONS) to Jobs in financial intermediation (ONS); M0 (IFS); M1 (BoE); M2 (BoE); M4 (BoE).

A.4. Consumer first-order conditions

$$\max_{c_t, x_t, l_{Gt}, l_{Qt}, s_{Gt}, q_t, d_t, k_{t+1}, h_{t+1}, M_{t+1}} \mathcal{V}(k_0, h_0, M_0; z_0, u_0, v_0) = E_0 \sum_{t=0}^{\infty} \beta^t \frac{(c_t x_t^\psi)^{1-\theta}}{1-\theta} \tag{28}$$

subject to

$$\lambda : w_t(l_{Gt} + l_{Qt})h_t + r_t s_{Gt} k_t + R_{Qt} d_t + \frac{T_t}{P_t} \geq \frac{P_{Qt}}{P_t} q_t + c_t + k_{t+1} - (1 - \delta_K) k_t + \frac{M_{t+1}}{P_t} - \frac{M_t}{P_t}, \tag{29}$$

$$\mu : \frac{M_t}{P_t} + \frac{T_t}{P_t} + q_t \geq c_t, \tag{30}$$

$$\varepsilon : c_t = d_t, \tag{31}$$

$$\psi : h_{t+1} = (1 - \delta_H) h_t + A_H [(1 - l_{Gt} - l_{Qt} - x_t) h_t]^\eta [(1 - s_{Gt}) k_t]^{1-\eta}, \tag{32}$$

$$0 = (c_t x_t^\psi)^{-\theta} x_t^\psi - \lambda_t - \mu_t + \varepsilon_t,$$

$$0 = \Psi c_t^{1-\theta} x_t^{\psi(1-\theta)-1} - \psi_t \eta A_H h_t (l_{Ht} h_t)^{\eta-1} (s_{Ht} k_t)^{1-\eta},$$

$$0 = \lambda_t w_t h_t - \psi_t \eta A_H h_t (l_{Ht} h_t)^{\eta-1} (s_{Ht} k_t)^{1-\eta},$$

$$0 = \lambda_t w_t h_t - \psi_t \eta A_H h_t (l_{Ht} h_t)^{\eta-1} (s_{Ht} k_t)^{1-\eta},$$

$$0 = \lambda_t r_t k_t - \psi_t (1 - \eta) A_H k_t (l_{Ht} h_t)^\eta (s_{Ht} k_t)^{-\eta},$$

$$0 = -\lambda_t \left(\frac{P_{Qt}}{P_t} \right) + \mu_t,$$

$$0 = \lambda_t R_{Qt} - \varepsilon_t,$$

$$0 = -\lambda_t + \beta E_t \{ \lambda_{t+1} [1 - \delta_K + r_{t+1} S_{G,t+1}] \} + \beta E_t \{ \psi_{t+1} (1 - \eta) A_H S_{Ht+1} (l_{Ht+1} h_{t+1})^\eta (S_{Ht+1} k_{t+1})^{1-\eta} \},$$

$$0 = -\psi_t + \beta E_t \{ \lambda_{t+1} W_{t+1} (l_{Gt+1} + l_{Qt+1}) \} + \beta E_t \{ \psi_{t+1} [1 - \delta_H + \eta A_H l_{Ht+1} (l_{Ht+1} h_{t+1})^{\eta-1} (S_{Ht+1} k_{t+1})^{1-\eta}] \},$$

$$0 = -\frac{\lambda_t}{P_t} + \beta E_t \left\{ \frac{\lambda_{t+1} + \mu_{t+1}}{P_{t+1}} \right\}.$$

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